



Multivariate Bayesian Forecasting under Functional Distortions in the χ^2 -metric

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The main spheres of Bayesian framework application:

- ▶ medical sciences;
- ▶ financial markets;
- ▶ bio-informatics.

Investigations on robustness of Bayesian forecasting:

- ▶ *Bayesian Robustness* / ed. by **G.O.Berger** et al., 1995;
- ▶ *Problemy odporności w byesowskiej analizie statystycznej* / **Marek Męczarski**, 1998;
- ▶ *Minimax Robustness of Bayesian Forecasting under Functional Distortions of Probability Densities* / **Alexey Kharin**, 2002.

Hypothetical Model of Bayesian Forecasting



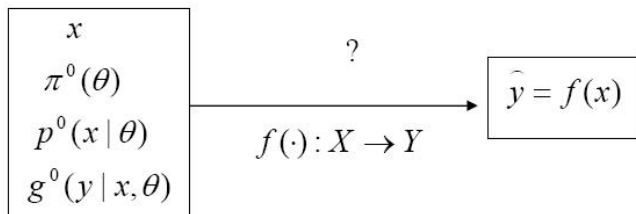
Let on a probability space (Ω, \mathcal{F}, P) be defined three random elements:

- ▶ the unobserved vector of model parameters

$$\theta \in \Theta \subseteq \mathbb{R}^m, \pi^0(\theta);$$

- ▶ the vector of observations $x = (x_t)_{t=1}^T \in X \subseteq \mathbb{R}^{n \times T}, p^0(x|\theta);$

- ▶ the value to be forecasted $y \in Y \subseteq \mathbb{R}^n, g^0(y|x, \theta).$



Hypothetical Risk Functional



The performance of a prediction statistics $f(\cdot) : X \rightarrow Y$ can be measured by **hypothetical risk functional** $r^0(f(\cdot))$:

$$r^0(f(\cdot)) = E_0 \{ \rho^2(f(x), y) \} = \int_X \int_Y s^0(x, y) \rho^2(f(x), y) dy dx, \quad (1)$$

$$s^0(x, y) = \int_{\Theta} g^0(y|x, \theta) p^0(x|\theta) \pi^0(\theta) d\theta. \quad (2)$$

$\rho(\cdot, \cdot)$ is the Euclidean distance function in \mathbb{R}^n .



Hypothetical Bayesian Prediction Density:

$$q^0(y|x) = \int_{\Theta} g^0(y|x, \theta) \pi^0(\theta|x) dx. \quad (3)$$

$$\pi^0(\theta|x) = \frac{p^0(x|\theta)\pi^0(\theta)}{p^0(x)}, p^0(x) = \int_{\Theta} p^0(x|\theta)\pi^0(\theta) d\theta. \quad (4)$$

Bayesian Prediction Statistics:

$$\hat{y} = f^0(x) = E_0 \{y|x\} = \int_Y y \cdot q^0(y|x) dy. \quad (5)$$

Bayesian Prediction Statistics is **optimal** with respect to the hypothetical risk functional.

Distortions of the Hypothetical Model



We consider functional distortions of priors, defined using χ^2 -metric:

$$\rho_{\chi^2}(h_1, h_2) = \int_U \frac{(h_1(u) - h_2(u))^2}{h_1(u)} du, \quad (6)$$

where p.d.f.s $h_1(u), h_2(u)$ are defined on U .

Suppose that θ is distributed according to an unknown p.d.f. $\pi^\varepsilon(\cdot) \in \Pi$:

$$\pi^0(\theta) \rightsquigarrow \pi^\varepsilon(\theta) \in \Pi = \{\Pi_\varepsilon : 0 \leq \varepsilon \leq \varepsilon_+\}, \quad (7)$$

$$\Pi_\varepsilon = \{\pi^\varepsilon(\cdot) : \rho_{\chi^2}(\pi^0(\cdot), \pi^\varepsilon(\cdot)) = \varepsilon_+^2\}.$$

Risk Functionals as Robustness Characteristics



We measure the performance of a prediction statistics

$f(\cdot) : X \rightarrow Y$ by the **risk functional** $r(\cdot, \cdot)$:

$$r(f(\cdot), \tilde{\pi}(\cdot)) = E \{ \rho^2(f(x), y) \} = \int_X \int_Y \tilde{s}(x, y) \rho^2(f(x), y) dx dy. \quad (8)$$

$$\tilde{s}(x, y) = \int_{\Theta} g^0(y|x, \theta) p^0(x|\theta) \tilde{\pi}(\theta) d\theta.$$

We use the **guaranteed upper risk functional** $r_*(\cdot)$ to measure the robustness of $f(\cdot)$:

$$r_*(f(\cdot)) = \sup_{\tilde{\pi}(\cdot) \in \Pi} r(f(\cdot), \tilde{\pi}(\cdot)). \quad (9)$$



We aim to find the **robust bayesian prediction statistics** $f_*(\cdot)$:

$$r_*(f_*(\cdot)) = \inf_{f(\cdot)} r_*(f(\cdot)). \quad (10)$$

The Conditional Risk Functional



As a Borelean function $\pi^\varepsilon(\cdot)$ from Π should be a p.d.f., the following ratios are valid:

$$\pi^\varepsilon(\theta) \geq 0, \theta \in \Theta, \int_{\Theta} \pi^\varepsilon(\theta) d\theta = 1.$$

The risk functional can be represented as:

$$r(f(\cdot); \pi^\varepsilon(\cdot)) = \int_{\Theta} \pi^\varepsilon(\theta) r_1(f(\cdot); \theta) d\theta, \quad (11)$$

$$r_1(f(\cdot); \theta) = \int_X \int_Y \rho^2(f(x), y) s^0(x, y|\theta) dy dx; \quad (12)$$

$$s^0(x, y|\theta) = g^0(y|x, \theta) p^0(x|\theta). \quad (13)$$



The **guaranteed upper risk functional** can be represented as

$$r_*(f(\cdot)) = \sup_{\pi^\varepsilon(\cdot) \in \Pi} r(f(\cdot); \pi^\varepsilon(\cdot)). \quad (14)$$

Denote **the conditional risk variance** as $\overset{\circ}{r}(f(\cdot); \theta)$:

$$\overset{\circ}{r}(f(\cdot); \theta) = r_1(f(\cdot); \theta) - E_0\{r_1(f(\cdot); \theta)\}. \quad (15)$$

Introduce **the critical distortions level**:

$$\varepsilon^*(f(\cdot)) = \frac{\sqrt{D_0\{r_1(f(\cdot); \theta)\}}}{\sup_{\theta \in \Theta} |\overset{\circ}{r}(f(\cdot); \theta)|}. \quad (16)$$



Theorem

Let the hypothetical forecasting model be distorted according to (7) and for any p.s. $f(\cdot) : X \rightarrow Y$ the distortion level $\varepsilon_+ \in [0, \varepsilon^*(f(\cdot))]$. Then the guaranteed upper risk functional can be represented as

$$r_*(f(\cdot)) = r(f(\cdot); \pi^*(\cdot)), \quad (17)$$

where the extreme p.d.f. $\pi^*(\cdot)$ is defined as

$$\pi^*(\theta) = \pi^0(\theta) \left(1 + \varepsilon_+ \frac{\overset{\circ}{r}(f(\cdot); \theta)}{\sqrt{D_0\{r_1(f(\cdot); \theta)\}}} \right). \quad (18)$$



Corollary

Under the theorem conditions the guaranteed upper risk can be represented as

$$r_*(f(\cdot)) = r_0(f(\cdot)) + \varepsilon_+ \sqrt{D_0\{r_1(f(\cdot); \theta)\}}, \quad (19)$$

where $r_0(f(\cdot))$ is the hypothetical risk functional:

$$r_0(f(\cdot)) = \int_{\Theta} r_1(f(\cdot); \theta) \pi^0(\theta) d\theta. \quad (20)$$

The Robust Prediction Statistics

Denote for $x \in X, y \in Y, \theta \in \Theta$:

$$F_\varepsilon(f(\cdot); x, y, \theta) = s^0(x, y|\theta) + \frac{\varepsilon (s^0(x, y|\theta) - s^0(x, y)) \overset{\circ}{r}(f(\cdot); \theta)}{\sqrt{D_0\{r_1(f(\cdot); \theta)\}}}. \quad (21)$$

$$\varepsilon^{**} = \inf_{f(\cdot)} \varepsilon^*(f(\cdot)). \quad (22)$$

Theorem

*Let the hypothetical forecasting model be distorted according to (7) and the distortion level $\varepsilon_+ \in [0, \varepsilon^{**}]$. Then the robust p.s. $f_*(\cdot)$ satisfies the following integral equation:*

$$f_*(x) = \frac{\iint_{Y\Theta} y \cdot \pi^0(\theta) F_{\varepsilon_+}(f_*(\cdot); x, y, \theta) d\theta dy}{\iint_{Y\Theta} \pi^0(\theta) F_{\varepsilon_+}(f_*(\cdot); x, y, \theta) d\theta dy}. \quad (23)$$





- ▶ The explicit expression (19) of the guaranteed upper risk allows calculating its deviation from the hypothetical risk for any p.s. $f(\cdot)$ and this deviation is at most of order $\mathcal{O}(\varepsilon_+)$.
- ▶ The integral equation (23) allows building iterative procedures for calculating the robust p.s. $f_*(\cdot)$:

$$f_{(0)} := f_0(x),$$

$$f_{(i)}(x) = \frac{\iint_{Y\Theta} y \cdot \pi^0(\theta) F_{\varepsilon_+}(f_{(i-1)}(\cdot); x, y, \theta) d\theta dy}{\iint_{Y\Theta} \pi^0(\theta) F_{\varepsilon_+}(f_{(i-1)}(\cdot); x, y, \theta) d\theta dy},$$

$$x \in X; i \in \mathbb{N}.$$



- ▶ **A.Kharin, P.Shlyk** *On Robustness of Multivariate Bayesian Forecasting under Functional Distortions of Priors*, BSU Bulletin, Minsk, vol.2, pp.103-107, 2006 (in russian)
- ▶ **A.Kharin, P.Shlyk** *On Robustness of Multivariate Bayesian Forecasting*, Proceedings for RobHD2004, Vorau, Austria, 2004.
- ▶ **A.Kharin** *Minimax Robustness of Bayesian Forecasting under Functional Distortions of Probability Densities*, Austrian Journal of Statistics, vol.31(2&3), pp.177-188, 2002.

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