Decomposing the Watson efficiency in a linear statistical model Jarkko Isotalo, Simo Puntanen Tampere, Finland Ka Lok Chu, George P. H. Styan McGill University, Montréal, Canada 1

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ABSTRACT

- In Chips1 (2004), we introduced a particular decomposition for the Watson efficiency of the OLS estimator
 \hfrac{\circ}{\beta}: product of three factors.
- Chips2 (2005) shows:

all three factors are related to the efficiencies of particular submodels or their *transformed versions*.

• There is:

an interesting connection between a particular reduction of the Watson efficiency and the concept of *linear sufficiency*. • The efficiency and specific *canonical correlations*.

• Decomposition for the

Bloomfield–Watson commutator criterion, and a condition for its specific reduction.

• Efficiency of $\mathbf{K}'\hat{\boldsymbol{\beta}}$



1. Introduction

(a serious one ...)

- 2. Decomposing the Watson efficiency
- 3. Canonical correlations
- 4. Decomposing the commutator criterion
- 5. References

1. Introduction

Consider the partitioned linear model

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon} = \mathbf{X}_1\boldsymbol{\beta}_1 + \mathbf{X}_2\boldsymbol{\beta}_2 + \boldsymbol{\varepsilon},$$
$$\mathscr{M}_{12} = \{\mathbf{y}, \, \mathbf{X}\boldsymbol{\beta}, \, \sigma^2 \mathbf{V}\},$$
$$\mathbf{E}(\mathbf{y}) = \mathbf{X}\boldsymbol{\beta}, \quad \mathbf{E}(\boldsymbol{\varepsilon}) = \mathbf{0}, \quad \operatorname{cov}(\mathbf{y}) = \operatorname{cov}(\boldsymbol{\varepsilon}) = \sigma^2 \mathbf{V}.$$

- **y** is an $n \times 1$ observable random vector,
- ε is an $n \times 1$ random error vector,
- **X** is a known $n \times p$ model matrix,
- β is a $p \times 1$ vector of unknown parameters.

$\mathbf{H} = \mathbf{P}_{\mathbf{X}}, \quad \mathbf{M} = \mathbf{I} - \mathbf{H}.$

The ordinary least squares estimator of $\mathbf{X}\boldsymbol{\beta}$:

$$OLSE(\mathbf{X}\boldsymbol{\beta}) = \mathbf{X}\hat{\boldsymbol{\beta}} = \hat{\mathbf{y}} = \mathbf{H}\mathbf{y} = \mathbf{P}_{\mathbf{X}}\mathbf{y}.$$

An unbiased estimator \mathbf{Gy} is the **best linear unbiased** estimator (BLUE) of $\mathbf{X\beta}$ if

$$\mathbf{GVG}' \leq \mathbf{BVB}' \quad \forall \mathbf{B} : \mathbf{BX} = \mathbf{X},$$

i.e.,

 $\mathbf{BVB'} - \mathbf{GVG'}$ is not $\forall \mathbf{B} : \mathbf{BX} = \mathbf{X}$.

In addition to \mathcal{M}_{12} , we consider

$$\begin{split} \mathscr{M}_1 &= \{\mathbf{y}, \, \mathbf{X}_1 \boldsymbol{\beta}_1, \, \mathbf{V}\}, \\ \mathscr{M}_{1\mathrm{H}} &= \{\mathbf{H}\mathbf{y}, \, \mathbf{X}_1 \boldsymbol{\beta}_1, \, \mathbf{H}\mathbf{V}\mathbf{H}\}, \\ \mathscr{M}_{12\cdot 1} &= \{\mathbf{M}_1 \mathbf{y}, \, \mathbf{M}_1 \mathbf{X}_2 \boldsymbol{\beta}_2, \, \mathbf{M}_1 \mathbf{V}\mathbf{M}_1\}. \end{split}$$

 $\mathcal{M}_1 = a \text{ small model}, \quad \mathcal{M}_{12 \cdot 1} = a \text{ reduced model}.$

- $\mathcal{M}_{12\cdot 1}$ is obtained by premultiplying \mathcal{M}_{12} by \mathbf{M}_1 .
- \mathcal{M}_{1H} is obtained by premultiplying \mathcal{M}_1 by **H**.

We consider a weakly singular model which means that \mathbf{V} may be singular but

$$\mathscr{C}(\mathbf{X}) \subset \mathscr{C}(\mathbf{V}).$$
 (*)

Under this model,

$$\begin{split} \tilde{\boldsymbol{\beta}} &= (\mathbf{X}'\mathbf{V}^{+}\mathbf{X})^{-1}\mathbf{X}'\mathbf{V}^{+}\mathbf{y},\\ \operatorname{cov}(\tilde{\boldsymbol{\beta}}) &= (\mathbf{X}'\mathbf{V}^{+}\mathbf{X})^{-1}\\ &= \mathbf{U}^{-1}[\mathbf{X}'\mathbf{V}\mathbf{X} - \mathbf{X}'\mathbf{V}\mathbf{M}(\mathbf{M}\mathbf{V}\mathbf{M})^{-}\mathbf{M}'\mathbf{V}\mathbf{X}]\mathbf{U}^{-1},\\ \mathrm{where} \ \mathbf{U} &= \mathbf{X}'\mathbf{X}. \end{split}$$

Hence,

$$\begin{split} \phi_{12} &= \operatorname{eff}(\hat{\boldsymbol{\beta}} \mid \mathscr{M}_{12}) = \frac{|\operatorname{cov}(\tilde{\boldsymbol{\beta}})|}{|\operatorname{cov}(\hat{\boldsymbol{\beta}})|} \\ &= \frac{|\mathbf{X}'\mathbf{X}|^2}{|\mathbf{X}'\mathbf{V}\mathbf{X}| \cdot |\mathbf{X}'\mathbf{V}^+\mathbf{X}|} \\ &= \frac{|\mathbf{X}'\mathbf{V}\mathbf{X} - \mathbf{X}'\mathbf{V}\mathbf{Z}(\mathbf{Z}'\mathbf{V}\mathbf{Z})^-\mathbf{Z}'\mathbf{V}\mathbf{X}|}{|\mathbf{X}'\mathbf{V}\mathbf{X}|} \\ &= |\mathbf{I}_p - \mathbf{X}'\mathbf{V}\mathbf{Z}(\mathbf{Z}'\mathbf{V}\mathbf{Z})^-\mathbf{Z}'\mathbf{V}\mathbf{X}(\mathbf{X}'\mathbf{V}\mathbf{X})^{-1}|, \end{split}$$

where $\mathbf{Z} = \mathbf{X}^{\perp}$.

Chips1 introduced a new decomposition for ϕ_{12} :

$$\operatorname{eff}(\hat{\boldsymbol{\beta}} \mid \mathscr{M}_{12}) = \operatorname{eff}(\hat{\boldsymbol{\beta}}_1 \mid \mathscr{M}_1) \cdot \operatorname{eff}(\hat{\boldsymbol{\beta}}_2 \mid \mathscr{M}_{12}) \cdot \alpha_1,$$

where $eff(\cdot | \cdot)$ refers to the Watson efficiency under a particular model, and α_1 is a specific determinant ratio.

- One of the key results in **Chips2**: $1/\alpha_1 =$ the efficiency of $\hat{\beta}_1$ under \mathcal{M}_{1H} .
- Another interesting observation: the reduction

$$\operatorname{eff}(\hat{\boldsymbol{\beta}} \mid \mathscr{M}_{12}) = \operatorname{eff}(\hat{\boldsymbol{\beta}}_2 \mid \mathscr{M}_{12}), \qquad (1.2)$$

is closely connected to the concept of linear sufficiency.

• Formally,

Fy is defined to be linearly sufficient for $X\beta$ under $\mathcal{M} = \{y, X\beta, V\}$ if

there exists a matrix \mathbf{A} such that \mathbf{AFy} is the **BLUE** of $\mathbf{X\beta}$.

... that was the the Introduction ...

I told you it was supposed to be a serious one ...

2. Decomposing the Efficiency

$$\operatorname{eff}(\hat{\boldsymbol{\beta}}_1 \mid \mathscr{M}_1) = \frac{|\mathbf{X}_1' \mathbf{X}_1|^2}{|\mathbf{X}_1' \mathbf{V} \mathbf{X}_1| \cdot |\mathbf{X}_1' \mathbf{V}^+ \mathbf{X}_1|} := \phi_{1/1},$$

 $\operatorname{eff}(\hat{\boldsymbol{\beta}}_2 \mid \mathscr{M}_{12}) = \operatorname{eff}(\hat{\boldsymbol{\beta}}_2 \mid \mathscr{M}_{12 \cdot 1})$

$$= \frac{|\mathbf{X}_{2}'\mathbf{M}_{1}\mathbf{X}_{2}|^{2}}{|\mathbf{X}_{2}'\mathbf{M}_{1}\mathbf{V}\mathbf{M}_{1}\mathbf{X}_{2}| \cdot |\mathbf{X}_{2}'\mathbf{\dot{M}}_{1}\mathbf{X}_{2}|}$$
$$:= \phi_{2/12},$$

where

$$\dot{\mathbf{M}}_1 = \mathbf{M}_1(\mathbf{M}_1\mathbf{V}\mathbf{M}_1)^{-}\mathbf{M}_1.$$

• Important Note:

The BLUE of β_2 under model \mathcal{M}_{12} coincides with the BLUE of β_2 under model $\mathcal{M}_{12\cdot 1}$, i.e.,

$$\tilde{\boldsymbol{\beta}}_2(\mathscr{M}_{12\cdot 1}) = \tilde{\boldsymbol{\beta}}_2(\mathscr{M}_{12}).$$

THEOREM 1 The total Watson efficiency ϕ_{12} of the OLSE(β) under the partitioned weakly singular linear model $\mathcal{M}_{12} = \{\mathbf{y}, \mathbf{X}\beta, \mathbf{V}\},\ can be expressed as the product$

$$\operatorname{eff}(\hat{\boldsymbol{\beta}} \mid \mathscr{M}_{12}) = \operatorname{eff}(\hat{\boldsymbol{\beta}}_1 \mid \mathscr{M}_1) \cdot \operatorname{eff}(\hat{\boldsymbol{\beta}}_2 \mid \mathscr{M}_{12}) \cdot \frac{1}{\operatorname{eff}(\hat{\boldsymbol{\beta}}_1 \mid \mathscr{M}_{1H})},$$

where

$$\operatorname{eff}(\hat{\boldsymbol{\beta}}_{1} \mid \mathcal{M}_{1H}) = \frac{|\mathbf{X}_{1}'\mathbf{X}_{1}|^{2}}{|\mathbf{X}_{1}'\mathbf{V}\mathbf{X}_{1}| \cdot |\mathbf{X}_{1}'(\mathbf{H}\mathbf{V}\mathbf{H})^{-}\mathbf{X}_{1}|},$$

$$\operatorname{eff}(\hat{\boldsymbol{\beta}}_1 \mid \mathscr{M}_{1\mathrm{H}}) := \phi_{1\mathrm{H}}$$

$$= \frac{|\mathbf{X}_1'\mathbf{X}_1|^2}{|\mathbf{X}_1'\mathbf{V}\mathbf{X}_1| \cdot |\mathbf{X}_1'(\mathbf{H}\mathbf{V}\mathbf{H})^{-}\mathbf{X}_1|}$$

$$= |\mathbf{I}_{p_1} - \mathbf{X}_1' \mathbf{V} \mathbf{M}_1 \mathbf{X}_2 (\mathbf{X}_2' \mathbf{M}_1 \mathbf{V} \mathbf{M}_1 \mathbf{X}_2)^{-1} \\ \cdot \mathbf{X}_2' \mathbf{M}_1 \mathbf{V} \mathbf{X}_1' (\mathbf{X}_1' \mathbf{V} \mathbf{X}_1)^{-1}|.$$

Next we take a look at the conditions for a particular **reduction** of the total Watson efficiency.

THEOREM 2 Let $\mathcal{M}_{12} = \{\mathbf{y}, \mathbf{X}\boldsymbol{\beta}, \mathbf{V}\}$ be a partitioned weakly singular linear model. Then the following statements are equivalent:

(a)
$$\operatorname{eff}(\hat{\boldsymbol{\beta}} \mid \mathscr{M}_{12}) = \operatorname{eff}(\hat{\boldsymbol{\beta}}_2 \mid \mathscr{M}_{12}),$$

(b) $\mathscr{C}(\mathbf{X}_1) \subset \mathscr{C}(\mathbf{V}\mathbf{X}),$

(c) **Hy** is linearly sufficient for $\mathbf{X}_1 \boldsymbol{\beta}_1$ under \mathcal{M}_1 .

3. The commutator criterion

THEOREM 3 The Bloomfield–Watson efficiency

$$\mathrm{eff}_{\mathrm{BW}}(\hat{\boldsymbol{\beta}} \mid \mathscr{M}_{12}) = \psi_{12} = \frac{1}{2} \|\mathbf{HV} - \mathbf{VH}\|^2 = \|\mathbf{HVM}\|^2$$

has the decomposition

$$\psi_{12} = \psi_{1/1} + \psi_{2/12} - \psi_{1\mathrm{H}};$$

$$\psi_{1/1} = \operatorname{eff}_{BW}(\hat{\boldsymbol{\beta}}_{1} \mid \mathscr{M}_{1}) = \|\mathbf{P}_{1}\mathbf{V}\mathbf{M}_{1}\|^{2},$$

$$\psi_{2/12} = \operatorname{eff}_{BW}(\hat{\boldsymbol{\beta}}_{2} \mid \mathscr{M}_{12}) = \|\mathbf{P}_{\mathbf{M}_{1}\mathbf{X}_{2}}\mathbf{V}\mathbf{M}\|^{2},$$

$$\psi_{1H} = \operatorname{eff}_{BW}(\hat{\boldsymbol{\beta}}_{1} \mid \mathscr{M}_{1H}) = \|\mathbf{P}_{1}\mathbf{V}\mathbf{P}_{\mathbf{M}_{1}\mathbf{X}_{2}}\|^{2}.$$

Consider the condition under which

• ψ_{12} reduces into $\psi_{2/12}$.

It is interesting that this condition differs from the corresponding condition regarding the Watson efficiency. THEOREM 4 The B-W efficiency has property

$$\psi_{12} = \psi_{2/12} \tag{3.2}$$

if and only if

$$\mathscr{C}(\mathbf{V}\mathbf{X}_1) \subset \mathscr{C}(\mathbf{X}). \tag{3.3}$$

Under a weakly singular linear model (3.3) becomes

$$\mathscr{C}(\mathbf{X}_1) \subset \mathscr{C}(\mathbf{V}^+\mathbf{X}).$$

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